

INSTITUTE OF ECONOMIC STUDIES, FACULTY OF SOCIAL SCIENCES CHARLES UNIVERSITY IN PRAGUE (established 1348)



ROBUST STATISTICS <sup>AND</sup> ECONOMETRICS

INSTITUTE OF ECONOMIC STUDIES FACULTY OF SOCIAL SCIENCES CHARLES UNIVERSITY IN PRAGUE

JAN ÁMOS VÍŠEK

Week 2

#### Content of lecture



Repetition of findings from previous lecture
How did we start to study the statistics ?

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How did we start to study the statistics ?



Main goals of robust statistics and problems to be solved

How did we start to study the statistics ?

### Content



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Repetition of findings from previous lecture
How did we start to study the statistics ?

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How did we start to study the statistics ?



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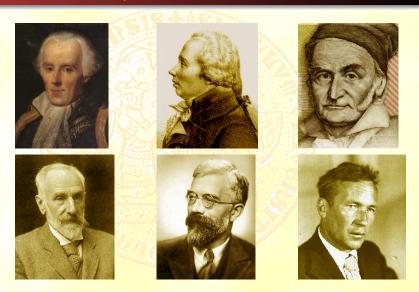
How did we start to study the statistics ?

# A small historical chimney-corner

# Could You assign some years or persons to this development ?



How did we start to study the statistics ?



How did we start to study the statistics ?



How did we start to study the statistics ?

### When did we start to study the statistics?



The first statistical society in the world



STATISTICAL SOCIETY in LONDON -

founded on February 21, 1834



the Czech Statistical Society

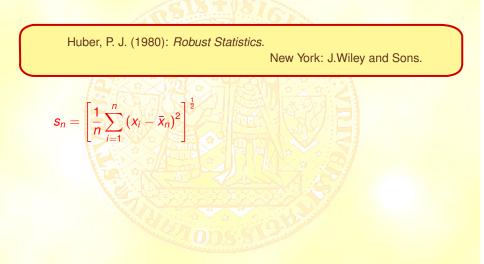
- Prague, March 29, 1990

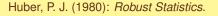
How did we start to study the statistics ?



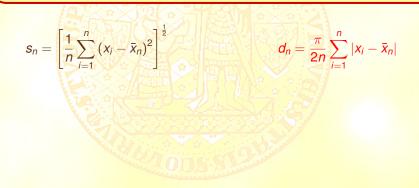
How did we start to study the statistics ?







New York: J.Wiley and Sons.



Huber, P. J. (1980): Robust Statistics.

New York: J.Wiley and Sons.

$$s_n = \left[\frac{1}{n}\sum_{i=1}^n (x_i - \bar{x}_n)^2\right]^{\frac{1}{2}} \qquad d_n = \frac{\pi}{2n}\sum_{i=1}^n |x_i - \bar{x}_n|$$
$$F(x) = (1 - \varepsilon)\Phi(x) + \varepsilon\Phi(\frac{x}{3})$$

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 $s_{n} = \left[\frac{1}{n}\sum_{i=1}^{n} (x_{i} - \bar{x}_{n})^{2}\right]^{\frac{1}{2}} \qquad d_{n} = \frac{\pi}{2n}\sum_{i=1}^{n} |x_{i} - \bar{x}_{n}|$   $F(x) = (1 - \varepsilon)\Phi(x) + \varepsilon\Phi(\frac{x}{3})$   $ARE_{F}(\varepsilon) = \lim_{n \to \infty} \frac{\operatorname{var}_{F} s_{n} / \mathbb{E}_{F}^{2} s_{n}}{\operatorname{var}_{F} d_{n} / \mathbb{E}_{F}^{2} d_{n}}$ 

How did we start to study the statistics ?

# Small deviation from exact model can cause ...

ε	0	0.001	0.002	0.05
$ARE(\varepsilon)$	0.876	0.948	1.016	2.035

How did we start to study the statistics ?

### Small deviation from exact model can cause ...

- PUIA	$\approx \times 1/26$	( C) TRON	1 a × a	13
<b>e b</b>	0	0.001	0.002	0.05
$ARE(\varepsilon)$	0.876	0.948	1.016	2.035

So, 5% of contamination  $\rightarrow d_n$  is two times better than  $s_n$ .

How did we start to study the statistics ?

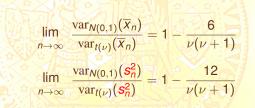
### Is 5% contamination too much or too little?

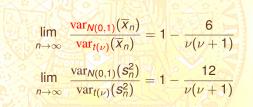
Hampel, F. R., E. M. Ronchetti, P. J. Rousseeuw, W. A. Stahel. (1986): *Robust Statistic - The Approach Based on Influence Curve.* New York: J.Wiley and Sons.

E. g. Switzerland has 6% of errors in mortality tables.

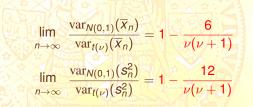
$$\lim_{n \to \infty} \frac{\operatorname{var}_{N(0,1)}(\overline{x}_n)}{\operatorname{var}_{t(\nu)}(\overline{x}_n)} = 1 - \frac{6}{\nu(\nu+1)}$$
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How did we start to study the statistics ?

# Is the efficiency really important or a bit misleading?

$\lim_{n \to \infty} \frac{\operatorname{var}_{N(0,1)}(T_n)}{\operatorname{var}_{t(\nu)}(T_n)}$	t9	<i>t</i> 5	t <sub>3</sub>
$\overline{x}_n$	0.93	0.80	0.50
s <sub>n</sub> <sup>2</sup>	0.83	0.40	0!

How did we start to study the statistics ?

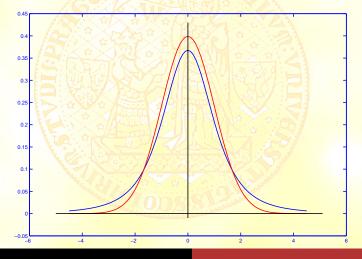
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How did we start to study the statistics ?

#### How far is Student density from the normal one ?

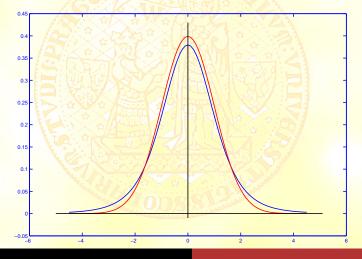
THE BLUE CURVE IS STANDARD NORMAL WHILE THE RED ONE IS THE STUDENT'S WITH 3 DEGREES OF FREEDOM.



How did we start to study the statistics ?

#### How far is Student density from the normal one ?

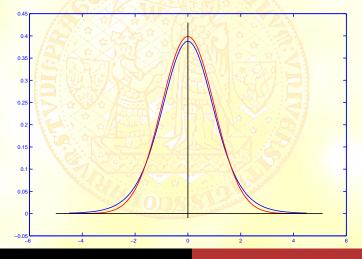
THE BLUE CURVE IS STANDARD NORMAL WHILE THE RED ONE IS THE STUDENT'S WITH 5 DEGREES OF FREEDOM.



How did we start to study the statistics ?

#### How far is Student density from the normal one ?

THE BLUE CURVE IS STANDARD NORMAL WHILE THE RED ONE IS THE STUDENT'S WITH 9 DEGREES OF FREEDOM.



A tacit hope in ingnoring deviations from ideal models was that they would not matter; that statistical procedures which were optimal under strict model would still be approximately optimal under the approximate model. Unfortunately, it turned out that this hope was often drastically wrong; even mild deviations often have much larger effects than were anticipated by most statisticians.

John W. Tukey (1960)

#### Let's study general reasons causing it - returning a few slides back.

Maximum likelihood - solving an extremal problem

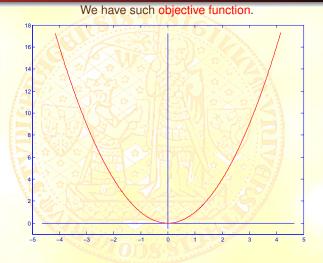
 $\hat{\theta}^{(ML,n)} = \underset{\theta \in \Theta}{\operatorname{arg\,max}} \prod_{i=1}^{n} f(x_i, \theta)$ 

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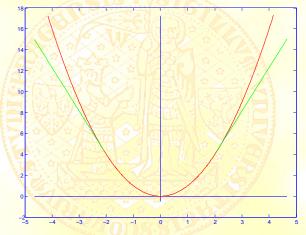
Let again  $f(x, \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} exp\left\{\frac{(x-\mu)^2}{2\sigma^2}\right\}$  and consider only  $\mu$ 

 $\Rightarrow \hat{\mu}^{(ML,n)} = \underset{\mu \in R}{\operatorname{arg\,min}} \left\{ \sum_{i=1}^{n} (x_i - \mu)^2 \right\}$ 

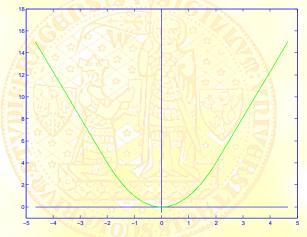
The observations with large  $(x_i - \mu)^2$ have a large influence on solution.

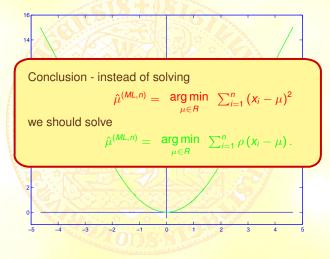






#### We should employ such objective function.





#### Let's study general reasons causing it - an alternative way.

Maximum likelihood - solving the normal equations

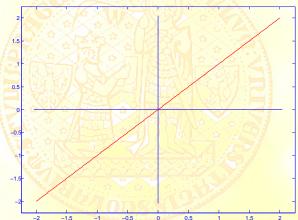
$$\hat{\theta}^{(ML,n)} = \underset{\theta \in \Theta}{\operatorname{arg\,max}} \prod_{i=1}^{n} f(x_i, \theta) = \underset{\theta \in \Theta}{\operatorname{arg\,max}} \sum_{i=1}^{n} \log \left( f(x_i, \theta) \right)$$

 $\hat{\theta}^{(ML,n)} = \arg_{\theta \in \Theta} \sum_{i=1}^{n} \frac{1}{f(x_i,\theta)} \cdot \frac{\partial f(x_i,\theta)}{\partial \theta} = 0$ 

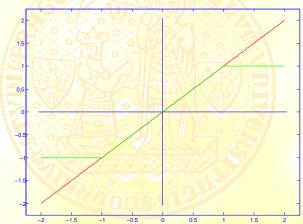
Let again  $f(x, \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} exp\left\{\frac{(x-\mu)^2}{2\sigma^2}\right\}$ , i.e.  $\frac{\partial f(x_i, \theta)}{\partial \mu} = f(x_i, \mu, \sigma^2) \cdot \frac{(x_i-\mu)}{\sigma^2}$ and consider only  $\mu \implies \hat{\mu}^{(ML,n)} = \underset{\mu \in R}{\operatorname{arg}} \left\{\sum_{i=1}^n (x_i - \mu) = 0\right\}$ 

The same conclusion:

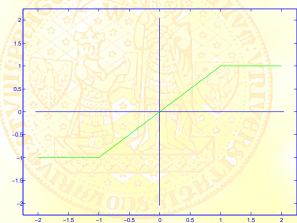
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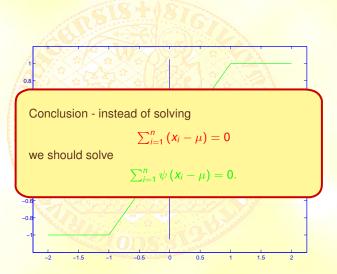
We have such influence function.



We should depress influence of large residuals.



We should employ such influence function.



How did we start to study the statistics ?

Motivation by historical experience

Ancient Egyptians and medieval French,

How did we start to study the statistics ?

Motivation by historical experience

- Ancient Egyptians and medieval French,
- Sir John William Rayleigh, Nobel Prize for Physics, 1904 (William Ramsay, Nobel Prize in chemistry, 1904)
   - 7 out of 15 atomic weight of "nitrogen" ⇒ argon,

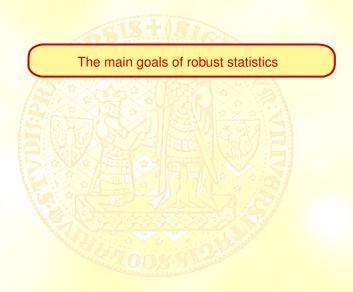
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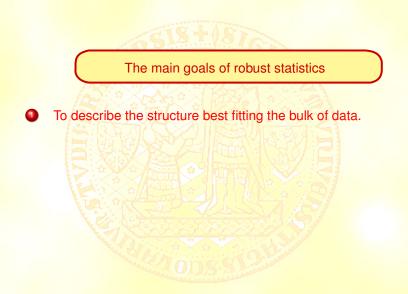
- Ancient Egyptians and medieval French,
- Sir John William Rayleigh, Nobel Prize for Physics, 1904 (William Ramsay, Nobel Prize in chemistry, 1904)
   - 7 out of 15 atomic weight of "nitrogen" ⇒ argon,
- J. B. Leon Foucalt 19. century, Albert Abraham Michelson - 1920 improved the method - 12 out of 16 measurements of light velocity.

(Remember Foucalt pendulum, 1851.)

#### Content







The main goals of robust statistics

- To describe the structure best fitting the bulk of data.
- To identify deviating data points (outliers) or deviating substructures for further treatment, if desired.

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The main goals of robust statistics

- To describe the structure best fitting the bulk of data.
- To identify deviating data points (outliers) or deviating substructures for further treatment, if desired.
- To identify and give a warning about highly influential data points (leverage points).
- To deal with unsuspected serial correlation, or more generally, with deviations from the assumed correlation structures.







- The occurrence of gross errors.
- 2 Rounding and grouping.

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- The model may have been conceived as an approximation anyway, e.g., by virtue of CLT.
- Apart of distributional assumptions, the assumption of independence (or of some specific correlation structure) may only be approximately fulfilled.

How have we attempted to cope with these tasks ?

Three approaches:

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Three approaches:

Huber's alternative to classical point estimation

via neighbourhoods.

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- 2 Huber's alternative to classical testing hypotheses via capacities.

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Three approaches:

- Huber's alternative to classical point estimation via neighbourhoods.
- Puber's alternative to classical testing hypotheses via capacities.
- Hampel's infinitesimal approach via Prokhorov metric and influence function.

# Huber's proposal to robustify point estimation

• Denote by  $\mathcal{H}$  the set of all distribution functions (d. f.'s).



### Huber's proposal to robustify point estimation

Denote by  $\mathcal{H}$  the set of all distribution functions (d. f.'s). 1 2 Select one fix  $F_{\theta_0} \in \mathcal{F}_{\Theta} = \{F_{\theta}\}_{\theta \in \Theta} \subset \mathcal{H}$ (called the parent or central distribution), fix also some  $H \in \mathcal{H}^* \subset \mathcal{H}$  and  $\delta > 0$ . Then put  $G_{\theta_0,\delta,H}(x) = (1-\delta)F_{\theta_0}(x) + \delta H(x).$ 

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Huber's proposal to robustify point estimation - an example

Previous lecture has recalled ML estimation 1  $\hat{\theta}^{(ML,n)} = \underset{\theta \in \Theta}{\operatorname{arg\,max}} \log \left\{ \prod_{i=1}^{n} f(x_{i},\theta) \right\},$ (1)

2

#### Huber's proposal to robustify point estimation - an example

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 $\hat{\theta}^{(ML,n)} = \arg \max_{\theta \in \Theta} \log \left\{ \prod_{i=1}^{n} f(x_i, \theta) \right\},$ (1) specifying example with  $f(x, \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} \exp \left\{ -\frac{(x-\mu)^2}{2\sigma^2} \right\}.$ 

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Specifying example with  $f(x, \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$ .

3 Putting for  $\Theta = (R \times R^+)$  and  $\theta = (\mu, \sigma^2)$  $\mathcal{F}_{\Theta} = \{N(x, \mu, \sigma^2)\}_{(\mu, \sigma^2) \in (R \times R^+)}$ 

we can write (1) as  $\hat{\theta}^{(ML,n)} = (\hat{\mu}^{(ML,n)}, \hat{\sigma}^{(ML,n)}) = \arg\max_{f \in \mathcal{F}_{\Theta}} \log \left\{ \prod_{i=1}^{n} f(x_i, \theta) \right\}.$ 

(The item 3 is rewritten on the next slide.)

#### Huber's proposal to robustify point estimation - an example

Putting for  $\Theta = (\mathbf{R} \times \mathbf{R}^+)$  and  $\theta = (\mu, \sigma^2)$ 0  $\mathcal{F}_{\Theta} = \left\{ N(x, \mu, \sigma^2) \right\}_{(\mu, \sigma^2) \in (B \times B^+)}$ we can write (1) as  $\hat{\theta}^{(ML,n)} = (\hat{\mu}^{(ML,n)}, \hat{\sigma}^{(ML,n)}) = \underset{f \in \mathcal{F}_{\Theta}}{\operatorname{arg\,max}} \log \left\{ \prod_{i=1}^{n} f(x_i, \theta) \right\}.$ 

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$$\hat{\theta}^{(ML,n,\varepsilon)} = (\hat{\mu}^{(ML,n,\varepsilon)}, \hat{\sigma}^{(ML,n,\varepsilon)}) = \arg\max_{g \in \mathcal{G}_{\Theta,\varepsilon}} \log \left\{ \prod_{i=1}^{n} g\left(x_{i}, \mu, \sigma^{2}, \delta, H\right) \right\}$$

where  $g(x_i, \mu, \sigma^2, \delta, H)$  is the density of

 $G_{\theta,\delta,H} = G_{\mu,\sigma^2,\delta,H} = (1 - \delta)\Phi(x,\mu,\sigma^2) + \delta H(x)$ and  $\mathcal{G}_{\Theta,\varepsilon,\mathcal{H}^*} = \{G_{\theta,\delta,H}\}_{\theta\in\Theta,\delta\leq\varepsilon,H\in\mathcal{H}^*}$ .

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For details see: Huber, P. J. (1964): Robust estimation of a location parameter. Ann. Math. Statist. 35, 73–101.

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Huber's proposal to robustify testing hypotheses

The proposal has the same character as the previous one but instead of considering the neighbourhoods of type

 $\mathcal{G}_{\Theta,\varepsilon,\mathcal{H}^*} = \{ \mathcal{G}_{\theta,\delta,H} \}_{\theta \in \Theta,\delta \leq \varepsilon,H \in \mathcal{H}^*} \cdot$ 

Peter Huber employed Choquet's capacities.

## Huber's proposal to robustify testing hypotheses

For details about the capacities see: Choquet, G. (1954): Theory of capacities. Annales de l'institut Fourier, 5 (1954), 131-295.

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A generalized Neyman-Pearson lemma

# Hampel's approach - a bit more mathematics



## Hampel's approach - a bit more mathematics

The Hampel's approach is based on two basic ideas and a nice fact:

0

The first idea - any estimator can be interpreted as a function T (say) from the space of all distribution functions  $\mathcal{H}$  to the parameter space  $\Theta$  (say).

## Hampel's approach - a bit more mathematics

- The first idea any estimator can be interpreted as a function T (say) from the space of all distribution functions H to the parameter space Θ (say).
- The second idea the function T can be studied by an infinitesimal calculus of limits, derivaties, integrals, etc.

# Hampel's approach - a bit more mathematics

- The first idea any estimator can be interpreted as a function T (say) from the space of all distribution functions  $\mathcal{H}$  to the parameter space  $\Theta$  (say).
- The second idea the function T can be studied by an infinitesimal calculus of limits, derivaties, integrals, etc.
- A nice fact the Kolmogorov-Smirnov result the empirical d.f. converge uniformly to the "true" underlying one.

# Hampel's approach - a bit more mathematics

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- The second idea the function T can be studied by an infinitesimal calculus of limits, derivaties, integrals, etc.
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The Hampel's approach is based on two basic ideas and a nice fact:

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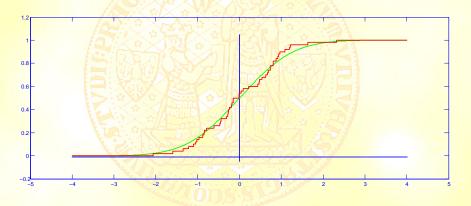
The first idea - any estimator can be interpreted as a function T (say) from the space of all distribution

Let's start - by an illustration - with the last topic, an exact mathematics will be delivered in some next lecture.

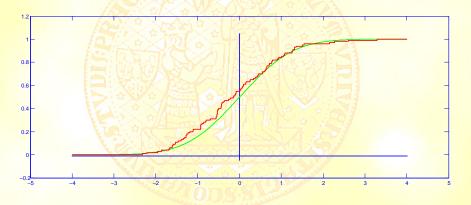
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Empirical distribution function - 50 observations.



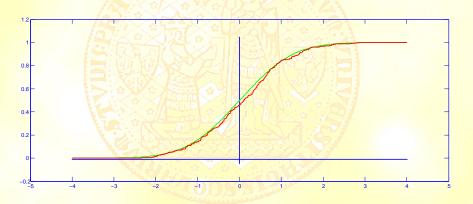
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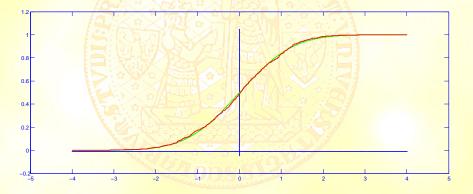
Empirical distribution function - 250 observations.



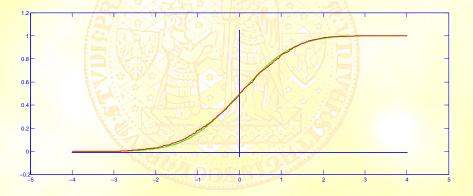
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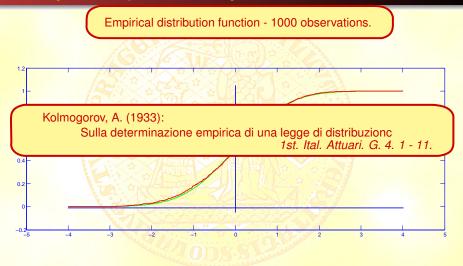


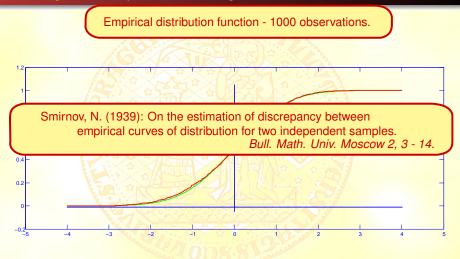
Empirical distribution function - 700 observations.

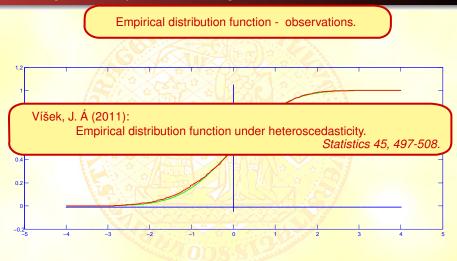


Empirical distribution function - 1000 observations.









Making preparation steps for explanation of Hampel's approach

Now, let us turn to the first idea:

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# Any estimator can be interpreted as a function T (say) from the space of all distribution functions $\mathcal{H}$ to the parameter space $\Theta$ (say).

Prior to it we need to recall something about the integration of functions.

### A preliminary intermezzo - the idea of integral

All of us learnt that the integral, say  $\int_a^b g(y) dy$ , is defined as follows:

Let  $a = y_0 < y_1 < y_2 < ... < y_n = b$  be an (equdistant) division of the interval [a, b] and for any  $i \in \{1, 2, ..., n\}$  let  $\tilde{y}_i \in [y_{i-1}, y_i]$ . Then put

$$\int_a^b g(y) \mathrm{d}y = \lim_{n \to \infty} \sum_{i=1}^n g(\tilde{y}_i)(y_i - y_{i-1}).$$

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We say that the integral is computed with respect to Lebesgue measure  $y_i - y_{i-1}$ - it is indicated by dy.

Let F be a d. f., f its density (in a general sense covering continuous as well as discrete r. v.'s) and Y random variable distributed according to F. Then the mean value of Y is given as

 $\boldsymbol{E}_{F}\boldsymbol{Y} = \int_{-\infty}^{\infty} \boldsymbol{y} f(\boldsymbol{y}) \, \mathrm{d}\boldsymbol{y} = \lim_{n \to \infty} \sum_{i=1}^{n} \tilde{y}_{i} f(\tilde{y}_{i})(\boldsymbol{y}_{i} - \boldsymbol{y}_{i-1}).$ 

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#### Notice the subindex in $E_F$

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$$EY = \lim_{n \to \infty} \sum_{i=1}^{n} \tilde{y}_i \left( F(y_i) - F(y_{i-1}) \right) = \int_{-\infty}^{\infty} y dF(y)$$

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where  $\tilde{y}_i$ 's are points at which is jump equal to  $\frac{1}{n}$ . Finally, let's recall that we have denoted the set of all d. f.'s by  $\mathcal{H}$ 

(in what follows we'll need it).

# The Hampel approach

Estimator as a function of distribution function

O Consider e. g.  $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$ .

Estimator as a function of distribution function

Consider e. g. x̄ = 1/n ∑<sub>i=1</sub><sup>n</sup> x<sub>i</sub>.
 Let F<sub>n</sub>(.) ∈ H be an empirical d. f. corresponding to the observations x<sub>1</sub>, x<sub>2</sub>, ..., x<sub>n</sub>, then T(F<sub>n</sub>) = ∫ xdF<sub>n</sub>(x) = 1/n ∑<sub>i=1</sub><sup>n</sup> x<sub>i</sub> = x̄ (because F<sub>n</sub>(x)has positive dF<sub>n</sub>(x) of size 1/n just at the points x<sub>1</sub>, x<sub>2</sub>, ..., x<sub>n</sub>).

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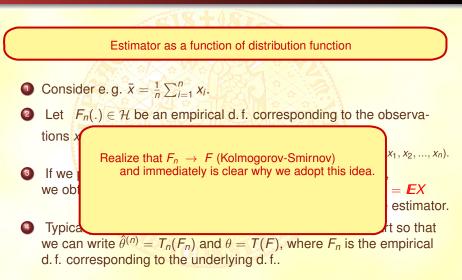
If we plug-in instead of empirical d. f. the underlying d. f. *F*, we obtain a function(al)  $T : \mathcal{H} \to \mathbb{R}^k$   $T(F) = \int x dF(x) = \mathbb{E}X$ which is a theoretical counterpart to the estimator.

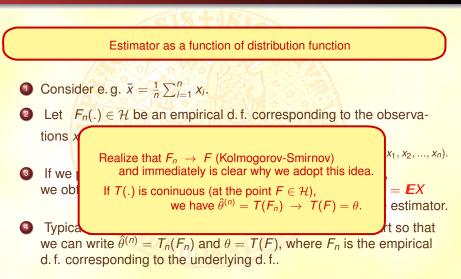
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Prior to it we need to carry out some preliminary explanation about the uncountably dimensional vector spaces.

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An example

$$\begin{bmatrix} \mathsf{B} & \longleftrightarrow & \mathsf{X} \\ \mathsf{C} & \longleftrightarrow & \mathsf{Y} \\ \mathsf{D} & \longleftrightarrow & \mathsf{Z} \end{bmatrix}$$

 $[A] \longleftrightarrow [W]$ 

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A \\
B \\
C \\
\hline
C \\
\hline
\end{array}$ 

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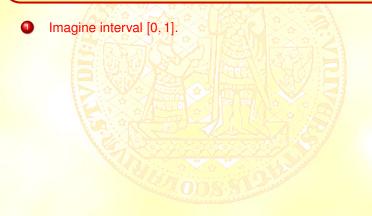
An example with infinite number of objects is on the next slide - consider the set of all positive integers

and the set of all positive rational numbers.

So, let's study cardinalities of the set of positive integers, say  $\mathcal{N},$  and of set of positive rational numbers,  $\mathcal{R}.$ 



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 Imagine interval [0, 1].
 Inside the interval [0, 1] there are at least all rationals of the type: <sup>1</sup>/<sub>2</sub>, <sup>1</sup>/<sub>3</sub>, <sup>2</sup>/<sub>3</sub>, <sup>1</sup>/<sub>4</sub>, <sup>3</sup>/<sub>4</sub>, <sup>1</sup>/<sub>5</sub>, <sup>2</sup>/<sub>5</sub>, <sup>3</sup>/<sub>5</sub>, <sup>4</sup>/<sub>5</sub>, etc. up to infinity.

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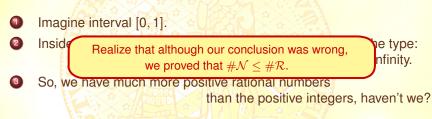
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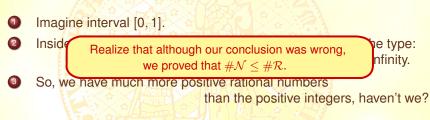
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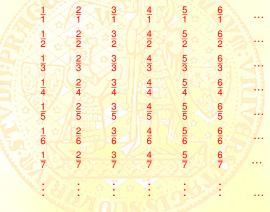
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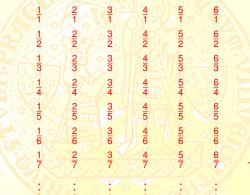
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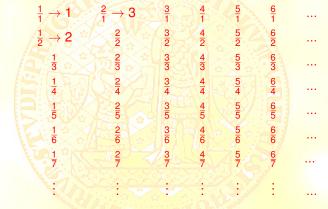
# Positive integers versus positive rational numbers

Starting the construction of mapping  $\mathcal{R}$  on  $\mathcal{N}$  - the first step:



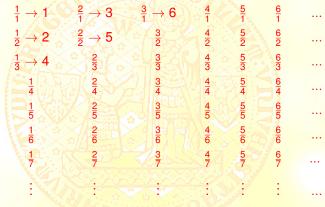
#### Positive integers versus positive rational numbers

# Starting the construction of mapping $\mathcal{R}$ on $\mathcal{N}$ - the second step:



#### Positive integers versus positive rational numbers

# Constructing the mapping $\mathcal{R}$ on $\mathcal{N}$ - the third step:



# Positive integers versus positive rational numbers

Constructing the mapping  $\mathcal{R}$  on  $\mathcal{N}$  - the fourth step :

$\frac{1}{1} \rightarrow 1$	$\frac{2}{1} \rightarrow 3$	$\frac{3}{1} \rightarrow 6$	$rac{4}{1}  ightarrow 10$	<u>5</u> 1	<u>6</u> 1	
$\frac{1}{2} \rightarrow 2$	$\frac{2}{2} \rightarrow 5$	$\frac{3}{2} \rightarrow 9$	4 2	5 2	<u>6</u> 2	
$\frac{1}{3} \rightarrow 4$	$\frac{2}{3} \rightarrow 8$	33	$\frac{4}{3}$	53	<u>6</u> 3	
$\frac{1}{4} \rightarrow 7$	2 4	34	4	5 4	<u>6</u> 4	
<u>1</u> 5	2 5	35	$\frac{4}{5}$	<u>5</u> 5	<u>6</u> 5	
<u>1</u> 6	26	<b>6 6</b>	$\frac{4}{6}$	<u>5</u> 6	<u>6</u> 6	
17	27	37	47	<u>5</u> 7	<u>6</u> 7	
	1900	650 6703	200	1	-	

# Positive integers versus positive rational numbers

Constructing the mapping  $\mathcal{R}$  on  $\mathcal{N}$  - the fourth step , etc.:

$\frac{1}{1} \rightarrow 1$	$\frac{2}{1} \rightarrow 3$	$\frac{3}{1} \rightarrow 6$	$\frac{4}{1} \rightarrow 10$	<u>5</u> 1	<u>6</u> 1	
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$\frac{1}{3} \rightarrow 4$	$\frac{2}{3} \rightarrow 8$	<u>3</u>	$\frac{4}{3}$	<u>5</u> 3	<u>6</u> 3	
$\frac{1}{4} \rightarrow 7$	24	34	4	<u>5</u> 4	<u>6</u> 4	
15	2 5	35	4 <u>5</u>	<u>5</u> 5	<u>6</u> 5	
$\frac{1}{6}$	2 6	3	$\frac{4}{6}$	<u>5</u> 6	<u>6</u> 6	
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	Sec.	Contraction of	2298	2/		
						Q.E.D.

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- We already know  $\mathcal{R}$  is countable.
- 2 If  $\mathcal{I}$  be countable  $\rightarrow \mathcal{RE}$  is countable.

## Cardinality of $\mathcal{R}$ (positive rationals), $\mathcal{I}$ (positive irrationals) and $\mathcal{RE}$ (positive

- We already know R is countable.
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- $\bullet \quad \text{However, we'll prove that } \mathcal{RE} \text{ is uncountable.}$

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Let's rewrite this scheme on the next slide.

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 $0. \ c_1^{(1)} \ c_2^{(1)} \ c_3^{(1)} \ c_4^{(1)} \ c_5^{(1)} \ c_6^{(1)} \ c_7^{(1)} \ c_8^{(1)} \dots$  $0. \ c_1^{(2)} \ c_2^{(2)} \ c_3^{(2)} \ c_4^{(2)} \ c_5^{(2)} \ c_6^{(2)} \ c_7^{(2)} \ c_8^{(2)} \ \dots$ 0.  $c_1^{(3)} c_2^{(3)} c_3^{(3)} c_4^{(3)} c_5^{(3)} c_6^{(3)} c_7^{(3)} c_8^{(3)} \dots$  $0. \ c_1^{(4)} \ c_2^{(4)} \ c_3^{(4)} \ c_4^{(4)} \ c_5^{(4)} \ c_6^{(4)} \ c_7^{(4)} \ c_8^{(1)} \ \dots$ 出入に人口間日

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Let's create a new real number (the upper index (*n*) indicates that it is "new" real number):

 $0. \ c_1^{(n)} \neq c_1^{(1)} \ c_2^{(n)} \neq c_2^{(2)} \ c_3^{(n)} \neq c_3^{(3)} \ c_4^{(n)} \neq c_4^{(4)} \ c_5^{(1)} \neq c_5^{(5)} \ c_6^{(1)} \neq c_6^{(6)} \ c_7^{(1)} \neq c_7^{(7)} \ c_8^{(1)} \neq c_8^{(8)} \ \dots$ 

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This new number does not coincide with any number in the sequence and it is a contradiction with the assumption that we had all real numbers in the sequence we studied above.

#### More shocking facts !!

We are going to prove a much more surprising result and what is nearly shocking - it can be done by trivial means.

## Cardinality (denote by #) of finite sets and sets of their subsets

Consider a finite set  $A = \{1, 2, ..., n\}$ , i. e. #A = n.

How much subsets it has (including the whole set A and the empty set)?



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Label of element	1 2 3	4	
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The set of all subsets, say A, has  $2^{\#A}$  elements ! So, #A < #A.

Does the last inequality hold also for infinite sets, i. e. is it still true that  $#A < 2^{#A}$ ?

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We are going to prove a much more surprising result and what is nearly shocking - it can be done by trivial means.

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• Consider any set A and all its subsets, say A.



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 $\forall (s \in A) \quad \exists (S \in A) \quad \text{so that} \quad S = \kappa(s).$ 

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It is a contradiction, Q.E.D.

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We proved that it holds generally that #A < #A.

Assume that  $s \in S$ . But point 3 then implies that  $s \notin S$ .

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- **Output** Consider any set A and all its subsets, say A.
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 $\forall (s \in A) \quad \exists (S \in A) \quad \text{so that} \quad S = \kappa(s).$ 

We proved that it holds generally that #A < #A. The set of real numbers is (more or less) the same thing as the set of all subsets of rational numbers.

- S Assume that  $\tilde{s} \in \tilde{S}$ . But point 3 then implies that  $\tilde{s} \notin \tilde{S}$ .
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# EXAMPLES:

- The sets of (all) rational numbers is countable.
- The sets of (all) irrational numbers is uncountable.
- The sets of (all) real numbers is uncountable.

Recalling the notion of vector space

Consider *p*-dimensional vector space, say U

then any vector  $u \in U$  has coordinates  $u_i, i \in \{1, 2, ..., p\}$ .

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Slightly generalize the notion of vector space

Consider countably-dimensional vector space, say Z
 → then any vector z ∈ Z has coordinates z<sub>i</sub>'s, i ∈ {1,2,...}.

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● Finally, consider uncountably-dimensional vector space, say  $\mathcal{F}$ → then any vector  $F \in \mathcal{F}$  has coordinates  $F(x), x \in R$ , say.

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We can again imagine that we have for any point of  $\mathcal{F}$  one mapping from the set R to the real line, i. e. for point F we have the mapping F(.) such that if we plug in some x from R, we obtain the x-th coordinate of F.

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## An example

In the previous lecture we met with the space of all distribution function  $\mathcal{H}$ . It is uncountably-dimensional vector space. Every d.f. *F*, including the empirical ones, is one point in it - convolution.



# THANKS FOR ATTENTION

